

2021 中级计量经济学作业 1 参考解答

1. 课本 28 页习题 3.3

(1) 正规方程组是 $(\mathbf{X}'\mathbf{X}\mathbf{b}) = \mathbf{X}'\mathbf{y}$, 也即 $\mathbf{X}'\mathbf{e} = \mathbf{0}$ 。其中 $\mathbf{X} = [\mathbf{1} \ \mathbf{x}_2 \ \dots \ \mathbf{x}_K]$, $\mathbf{x}_k = [x_{1k} \ \dots \ x_{nk}]'$ 。因此 $\sum_{i=1}^n e_i = 0$, $\sum_{i=1}^n x_{ik}e_i = 0$ 。

(2) $\frac{1}{n} \sum_{i=1}^n \hat{y}_i = \frac{1}{n} \sum_{i=1}^n (y_i - e_i) = \frac{1}{n} \sum_{i=1}^n y_i$, 因为 $\sum_{i=1}^n e_i = 0$ 。

(3)

$$\begin{aligned} (\mathbf{y} - \bar{\mathbf{y}})'(\mathbf{y} - \bar{\mathbf{y}}) &= (\mathbf{y} - \hat{\mathbf{y}} + \hat{\mathbf{y}} - \bar{\mathbf{y}})'(\mathbf{y} - \hat{\mathbf{y}} + \hat{\mathbf{y}} - \bar{\mathbf{y}}) \\ &= \mathbf{e}'\mathbf{e} + 2(\hat{\mathbf{y}} - \bar{\mathbf{y}})'\mathbf{e} + (\hat{\mathbf{y}} - \bar{\mathbf{y}})'(\hat{\mathbf{y}} - \bar{\mathbf{y}}) \end{aligned} \quad (1)$$

(2)

其中, $\hat{\mathbf{y}}'\mathbf{e} = (\mathbf{X}\mathbf{b})'\mathbf{e} = \mathbf{b}'\mathbf{X}'\mathbf{e} = \mathbf{0}$, $\bar{\mathbf{y}}'\mathbf{e} = \bar{y}'(\sum_{i=1}^n e_i) = 0$ 。因此

$$(\mathbf{y} - \bar{\mathbf{y}})'(\mathbf{y} - \bar{\mathbf{y}}) = \mathbf{e}'\mathbf{e} + (\hat{\mathbf{y}} - \bar{\mathbf{y}})'(\hat{\mathbf{y}} - \bar{\mathbf{y}}) \quad (3)$$

2. 课本 28 页习题 3.4

写成向量形式:

$$[\text{Corr}(y_i, \hat{y}_i)]^2 = \frac{(\mathbf{y} - \bar{\mathbf{y}})'(\hat{\mathbf{y}} - \bar{\mathbf{y}})(\mathbf{y} - \bar{\mathbf{y}})'(\hat{\mathbf{y}} - \bar{\mathbf{y}})}{(\mathbf{y} - \bar{\mathbf{y}})'(\mathbf{y} - \bar{\mathbf{y}})(\hat{\mathbf{y}} - \bar{\mathbf{y}})'(\hat{\mathbf{y}} - \bar{\mathbf{y}})} \quad (4)$$

$$R^2 = \frac{(\hat{\mathbf{y}} - \bar{\mathbf{y}})'(\hat{\mathbf{y}} - \bar{\mathbf{y}})}{(\mathbf{y} - \bar{\mathbf{y}})'(\mathbf{y} - \bar{\mathbf{y}})} = \frac{(\hat{\mathbf{y}} - \bar{\mathbf{y}})'(\hat{\mathbf{y}} - \bar{\mathbf{y}})(\hat{\mathbf{y}} - \bar{\mathbf{y}})'(\hat{\mathbf{y}} - \bar{\mathbf{y}})}{(\mathbf{y} - \bar{\mathbf{y}})'(\mathbf{y} - \bar{\mathbf{y}})(\hat{\mathbf{y}} - \bar{\mathbf{y}})'(\hat{\mathbf{y}} - \bar{\mathbf{y}})} \quad (5)$$

(4) 和 (5) 分母相同, 仅需证明分子相同即可。注意到 $\mathbf{y} - \bar{\mathbf{y}} = \hat{\mathbf{y}} - \bar{\mathbf{y}} + \mathbf{e}$, (4) 式分子可改写为

$$(\hat{\mathbf{y}} - \bar{\mathbf{y}} + \mathbf{e})'(\hat{\mathbf{y}} - \bar{\mathbf{y}})(\hat{\mathbf{y}} - \bar{\mathbf{y}} + \mathbf{e})'(\hat{\mathbf{y}} - \bar{\mathbf{y}}) \quad (6)$$

而 $\hat{\mathbf{y}}'\mathbf{e} = \bar{\mathbf{y}}'\mathbf{e} = 0$, 因此 (4) 和 (5) 分子相同。

3. 课本 63 页习题 5.2

$$(\hat{\beta}_n - \beta) = \frac{1}{\sqrt{n}} \sqrt{n}(\hat{\beta}_n - \beta) \quad (7)$$

其中 $\text{plim}_{n \rightarrow \infty} \frac{1}{\sqrt{n}} = 0$ 。由 Slutsky's theorem, $\hat{\beta}_n \xrightarrow{p} \beta$ 。

4. 课本 63 页习题 5.4

不一定。例如, $g_t = \varepsilon_t$, 其中 ε_t 独立, $\varepsilon_t \sim N(0, t)$ 。则 $\{g_t\}$ 是鞅差分过程, 但不是白噪声。

5. 课本 63 页习题 5.5

令 $\text{Var}(\varepsilon_i) = \sigma^2$ 。

(1) $\{x_i \varepsilon_i\}$ 独立但不同分布。例如, $\text{Var}(x_1 \varepsilon_1) = x_1^2 \sigma^2$, $\text{Var}(x_2 \varepsilon_2) = x_2^2 \sigma^2$ 。

(2) 不存在自相关。 $\text{Cov}(x_i \varepsilon_i, x_j \varepsilon_j) = x_i x_j \text{Cov}(\varepsilon_i, \varepsilon_j) = 0$

(3) 是鞅差分序列。 $E(x_i \varepsilon_i | x_{i-1} \varepsilon_{i-1}, \dots, x_1 \varepsilon_1) = 0$ 。

(4) 不是严格平稳序列。例如, $x_1 \varepsilon_1$ 和 $x_2 \varepsilon_2$ 的分布不相同。

6. 证明“线性假设” ($y_i = \mathbf{x}'_i \boldsymbol{\beta} + \epsilon_i$) 和“严格外生性假设” ($E(\epsilon_i | \mathbf{X}) = 0$) 同时满足时, “条件期望函数” (Conditional Expectation Function, CEF) 是线性的, 也即

$$E[y_i | \mathbf{X}] = \mathbf{x}'_i \boldsymbol{\beta} \quad (i = 1, 2, \dots, n). \quad (8)$$

反之, 证明上式成立时, 存在 ϵ 使“线性假设”和“严格外生性假设”成立。

证明:

必要性:

$$E[y_i | \mathbf{X}] = E[\mathbf{x}'_i \boldsymbol{\beta} + \epsilon_i | \mathbf{X}] = E[\mathbf{x}'_i \boldsymbol{\beta} | \mathbf{X}] + E[\epsilon_i | \mathbf{X}] = \mathbf{x}'_i \boldsymbol{\beta}$$

充分性: y_i 可写成 $y_i = E[y_i | \mathbf{X}] + y_i - E[y_i | \mathbf{X}] = \mathbf{x}'_i \boldsymbol{\beta} + \epsilon_i$, 其中 $\epsilon_i \equiv y_i - E[y_i | \mathbf{X}]$, $E[\epsilon_i | \mathbf{X}] = 0$ 。

7. (请附上代码以及运算结果) 运用自己熟悉的软件 (R, Stata, Matlab 等) 生成一组随机数 $\mathbf{y}, \mathbf{x}_1, \mathbf{x}_2$, 用 OLS 估计 $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \epsilon$ (此步骤可以用软件自动完成)。用软件的矩阵运算, 计算:
1. 同方差假设下参数 β_1, β_2 估计的标准误 (standard error)
 2. 异方差稳健标准误 (robust standard error)
 3. 将手动计算的结果与软件的计算结果进行对比, 检查数值是否相等。

以 R 代码为例:

```
library(data.table)
set.seed(42)
n <- 100
K <- 3
y <- rnorm(n)
x1 <- rnorm(n, 3, 4)
x2 <- rnorm(n, 10, 10)
dt <- data.table(y=y, x1=x1, x2=x2)
head(dt)

##           y           x1           x2
## 1:  1.3709584  7.8038615 -10.009292
## 2: -0.5646982  7.1790043  13.337772
## 3:  0.3631284 -1.0128346  21.713251
## 4:  0.6328626 10.3939276  30.595392
## 5:  0.4042683  0.3329064  -3.768616
## 6: -0.1061245  3.4220552  -1.508556

# Regression
reg <- lm(y~x1+x2, data=dt)
e <- reg$residuals
# independent variable matrix
X <- as.matrix(dt[,.(1,x1,x2)])
```

1. 同方差假设的标准误

手动计算

```
s_squared <- as.vector((t(e) %*% e)/(n-K))
# S_XX --> E[xx']
S_XX <- (t(X) %*% X) /n
S_XX_inverse <- solve(S_XX)
```

```
var_homo <- (s_squared * S_XX_inverse)/n
# standard error under homoskedasticity
beta1hat_std_homo <- sqrt(var_homo[2,2])
beta2hat_std_homo <- sqrt(var_homo[3,3])
cat("hat(beta1) 的标准误, 矩阵计算: ", beta1hat_std_homo)
```

```
## hat(beta1) 的标准误, 矩阵计算: 0.02897481
```

```
cat("hat(beta2) 的标准误, 矩阵计算: ", beta2hat_std_homo)
```

```
## hat(beta2) 的标准误, 矩阵计算: 0.01030401
```

软件计算

软件给出的标准误: $\hat{\beta}_1$ 对应 x_1 的 Std. Error, $\hat{\beta}_2$ 对应 x_2 的 Std. Error。

```
summary(reg)
```

```
##
## Call:
## lm(formula = y ~ x1 + x2, data = dt)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -3.15652 -0.60222  0.07537  0.66410  2.55446
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.15034     0.16123   0.932   0.353
## x1           0.01204     0.02897   0.415   0.679
## x2          -0.01513     0.01030  -1.468   0.145
##
## Residual standard error: 1.04 on 97 degrees of freedom
## Multiple R-squared:  0.0227, Adjusted R-squared:  0.002549
## F-statistic: 1.126 on 2 and 97 DF,  p-value: 0.3284
```

2. 异方差稳健的标准误

手动计算

```
# S = E[epsilon_i^2 x_i x_i']
S_hat <- (t(X) %*% diag(e^2) %*% X)/n
var_hetero <- (S_XX_inverse %*% S_hat %*% S_XX_inverse)/n
# standard error under heteroskedasticity
beta1hat_std_hetero <- sqrt(var_hetero[2,2])
beta2hat_std_hetero <- sqrt(var_hetero[3,3])
cat("hat(beta1) 的稳健标准误, 矩阵计算: ", beta1hat_std_hetero)
```

```
## hat(beta1) 的稳健标准误, 矩阵计算: 0.03416057
```

```
cat("hat(beta2) 的稳健标准误, 矩阵计算: ", beta2hat_std_hetero)
```

```
## hat(beta2) 的稳健标准误, 矩阵计算: 0.01047688
```

软件计算

```
library(lmtest)
```

```
## Loading required package: zoo
```

```
##
```

```
## Attaching package: 'zoo'
```

```
## The following objects are masked from 'package:base':
```

```
##
```

```
##      as.Date, as.Date.numeric
```

```
library(sandwich)
```

```
# Stata's robust standard error is slightly different from R's.
```

```
# The n-K adjustment is applied in Stata. If you want to reproduce the results in Stata,
```

```
# use type="HC1" in below.
```

```
coefTest(reg, vcov = vcovHC(reg, type="HCO"))
```

```
##
```

```
## t test of coefficients:
```

```
##
```

```
##           Estimate Std. Error t value Pr(>|t|)
```

```
## (Intercept)  0.150340  0.174230  0.8629  0.3903
```

```
## x1           0.012037  0.034161  0.3524  0.7253
```

```
## x2          -0.015129  0.010477 -1.4441  0.1519
```